Elevated Sleep Disturbance among Blacks in an Urban Family Medicine Practice

Wilfred R. Pigeon, PhD, Kathi Heffner, PhD, Paul Duberstein, PhD, Kevin Fiscella, MD, MPH, Jan Moynihan, PhD, and Benjamin P. Chapman, PhD

Purpose: Blacks experience a number of health disparities. Sleep disturbances contribute to poor health. This preliminary study explores whether a disparity in sleep disturbances exists among blacks compared with whites and others.

Methods: A cross-sectional study was conducted in a sample (n = 92) of urban primary care patients (52% black, 46% white, and 2% other) from a university-based family medicine practice. Mean (SD) age was 51.9 years (8.9 years). Participants completed the Pittsburgh Sleep Quality Index, the Center for Epidemiologic Studies Depression Scale, Revised, and a checklist of chronic health conditions.

Results: The rate of clinically meaningful sleep disturbance was 71%. In bivariate logistic regressions, black race was associated with sleep disturbance (odds ratio [OR], 3.00; 95% CI, 1.17–7.69). Controlling for income attenuated that association by about 11% (race OR, 2.71; 95% CI, 1.04–7.06). Education explained about 35% (race OR, 2.39; 95% CI, 0.89–6.42). Adjustment for depression, chronic illness, and education simultaneously resulted in an estimate for race of OR, 2.44; 95% CI, 0.85–7.01.

Conclusion: Being black is associated with a sleep disturbance that is accounted for only partially by depression, socioeconomic status, and disease burden. Black primary care patients may benefit from additional screening and monitoring of sleep difficulties. (J Am Board Fam Med 2011;24:161–168.)

Keywords: Blacks, Health Care Disparities, Insomnia, Minority Health, Sleep Disorders

Regardless of economic status, blacks in the United States experience chronic stress and consequent poor health because of social disadvantages and stressor-related allostatic load. We use the term “black” to refer to people who self-identify as being of African descent, a very diverse group that includes US-born African Americans whose lineages have endured slavery and its consequences, people of Afro-Caribbean descent, and recent immigrants from Africa. Sleep disturbance is an often unrecognized and seldom examined component of illness that contributes to health status. Poor sleep may represent a response to stress, a component contributing to allostatic load, and/or a mediator between psychological distress and neuroendocrine effects, leading to negative health outcomes. Thus, sleep disturbance may constitute an important contributor to health disparities among blacks.

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Sleep disturbances are associated with considerable morbidity, and increased mortality, and even susceptibility to developing a common cold. The most common form of sleep disturbance, chronic insomnia, occurs in approximately 10% of the general population and is associated with substantial economic costs and health consequences, including higher health care expenses, with direct costs estimated at US $13 billion per annum in physician visits, prescriptions, and procedures. Not surprisingly, sleep disturbances, particularly insomnia, are highly prevalent among primary care patients. Chronic insomnia is an independent risk factor for hypertension, which is a key driver of the disproportionate overall illness burden in socioeconomically disadvantaged blacks.

In general, both subjective and objective sleep disturbances are more pronounced in individuals with lower socioeconomic status. Likewise, greater subjective and objective sleep disturbances have been observed among blacks compared with whites, with some exceptions. Findings have consistently shown that blacks have higher rates of sleep apnea, shorter or longer habitual sleep durations, and more objective indicators of insomnia than whites. Subjectively, blacks report insomnia or indicators of insomnia at similar or lower rates than whites, perhaps because of under-endorsement of insomnia in some samples of black patients. Taken together, the available literature suggests a disparity with respect to sleep disturbances among black populations.

In this article we focus specifically on whether blacks have more sleep disturbances (measured with a validated sleep instrument) than whites and others in a clinical sample. We targeted middle-aged and older adults, given the cumulative nature of chronic illness and the increase in sleep disturbances across the life span. We recruited participants from an urban primary care setting because of the increased burden of chronic disease among socioeconomically disadvantaged urbanites. We hypothesized that blacks would be more likely to report sleep disturbances.

**Methods**

**Participants and Procedures**

Patients aged 40 years and older were recruited in person and by flyers posted in the waiting room at the Family Medicine Center of the University of Rochester Medical Center, directing them to call a research coordinator to schedule a research appointment at the Family Medicine Center or the General Clinical Research Center (n = 107). The flyer and the informed consent document indicated that the time commitment for the study would be approximately 3 hours and that participants would be compensated $50 for their participation. At the appointment, after providing written informed consent, participants completed an interview assessing demographics and psychosocial circumstances, the Pittsburgh Sleep Quality Index (PSQI), the Center for Epidemiologic Studies Depression Scale, Revised (CESD-R), and a checklist of chronic health conditions (see “Instruments,” below), as well as a blood draw (these data have been reported elsewhere). Fifteen participants did not complete the PSQI; they did not differ from the rest of the sample on any demographic variables.

Respondents completing all study instruments (n = 92) were, on average, middle-aged (mean, 51.9 years; SD, 8.91 years) and the majority were women (77.2%). They had an income level of less than $20,000 per year (62%), were not currently married (73.9%), and had children (83.7%). A little more than one quarter of the sample did not complete high school (26.1%). Participants were primarily black (52.2%) and white (46.6%), with one participant each endorsing only American Indian/Native and other, respectively. For analyses, the latter 2 participants (2.2%) were combined with whites into one “white/other” group. The racial diversity of the sample mirrored the diversity of the entire clinic population. Additional racial and other characteristics of the sample are further detailed in Table 1.

**Instruments**

**PSQI**

The PSQI is a 24-item scale that measures sleep disturbances along 7 dimensions. Scores from these 7 dimensions (ranging from 0–3) are individually reported as component scores and summed to derive a global sleep quality score (0–21); a score of >5 has been demonstrated to be indicative of a sleep disturbance. A subsequent reliability and validation study comparing the PSQI with sleep diaries and with objective polysomnographic measures of sleep among patients with insomnia found that using a cutoff of >6 resulted in the best sensitivity and specificity with respect to insomnia, a convention that we adopted in this study. Cronbach’s α internal consistency in the
The current sample was 0.78, which compares with the 0.83 reported in the original validation study. We report the global score and each of the component scores, as well as 2 single items from the scale, that are useful self-report measures of sleep continuity disturbance observed in insomnia samples. These were total sleep duration (dichotomized at 6.5 hours) and sleep efficiency, which is the ratio of time spent in bed to total sleep duration (dichotomized at 85%).

**CESD-R**

The CESD-R is a 20-item measure of depressive symptoms during the previous week. Responses involve a 4-point Likert scale ranging from 0 (“not at all”) to 3 (“nearly every day”). Cronbach’s α internal consistency in the current sample was 0.93. For descriptive purposes we report both the total CESD-R scores and total score with the sleep items removed, though only the latter were used in analyses.

### Other Health Factors

Patients completed a checklist of chronic conditions adapted from that used in the Midlife Development in the United States survey. The self-report checklist asks, Has a doctor ever told you you’ve had any of the following problems? Place a number to the left of each problem below: 0 = never; 1 = previously had this; 2 = have this now or within the last year. The 25 common chronic med-
ical conditions span respiratory, gastrointestinal, neurological, endocrine, and cardiovascular problems; conditions include “asthma,” “recurring backache,” “diabetes,” and “persistent trouble with your mouth or gums.” All items rated a 2 were summed to form a morbidity index. Body weight was also available from primary care charts for approximately 90% of the sample.

**Statistical Analysis**

After descriptive statistics, a series of bivariate logistic regressions were conducted, with sleep disturbance (PSQI >6), sleep duration (<6.5 hours per night), and sleep efficiency (<85%) as the dependent variables. The main predictor of interest was race (black vs white/other); additional predictors were sex, employment status (employed versus unemployed), annual household income (<$20,000 vs ≥$20,000 per year), education level (no high school diploma vs high school diploma vs a college degree), depressive symptoms (by CESD-R quartile with sleep items removed), and disease burden (Morbidity Index). We also considered age and body weight as potential covariates. We tested each covariate separately in models predicting PSQI >6 from race, and subsequently included all significant covariates in the final model. To examine if any of the other predictors explained observed associations between race and sleep disturbance, we computed the change in estimate that resulted from including each variable in the model ([unadjusted odds ratio (OR) for race – adjusted OR for race] / [unadjusted OR for race − 1]). Although neither the CESD-R category absent sleep items nor the Morbidity Index differed by race or by income, we conservatively adjusted for depression and common medical conditions. Finally, to further assess whether specific chronic illnesses were associated with sleep disturbance, contingency analyses were conducted for the presence or absence of specific conditions derived from the self-reported checklist and PSQI scores above or below our cutoff. All analyses were conducted using SPSS software version 17.0 (SPSS, Inc., Chicago, IL).

**Results**

**Descriptive Findings**

Descriptive statistics for the global PSQI, the 7 PSQI component scores, the sleep duration and efficiency items, the depressive symptomatology score from the CESD-R (excluding the sleep items), age, body weight, and morbidity are presented in Table 2. Using a cutoff of 6 to identify the presence of clinical sleep disturbance or insomnia, 70.7% of the sample (n = 65) scored above the cutoff. The high mean (sd) PSQI global score of 10.0 (4.9) is elevated compared with a healthy sample and a sample with suspected sleep apnea, and it is comparable to values observed in an insomnia sample from the original PSQI validation studies.

Using the CESD-R (including the sleep items) to classify individuals who have a moderate to high likelihood of major depression, 46.7% (n = 43) met the standard CESD-R cutoff of 16; 33.7% (n = 31) met the higher suggested cutoff of 21 for older adults. As might be expected, Pearson correlation coefficients showed that the CESD-R score (excluding the sleep items) was correlated with poorer global sleep quality (r = 0.53; P < .001), shorter sleep duration (r = 0.24; P < .05), and lower sleep efficiency (r = 0.23; P < .05). Patients with and without sleep disturbances did not significantly differ by age (PSQI global sleep quality score >6 mean, 51.14; SD, 7.71; PSQI score ≤6 mean, 53.59; SD, 11.28; P = .23) or body weight (PSQI global sleep quality score >6 mean, 201.23; SD, 42.78; PSQI score ≤6 mean, 196.22; SD, 46.17; P = .62). The differences in Morbidity Index scores for patients with and without sleep disturbances

<table>
<thead>
<tr>
<th>Variable</th>
<th>Mean (SD)</th>
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<tbody>
<tr>
<td>PSQI global sleep quality</td>
<td>10.00 (4.88)</td>
</tr>
<tr>
<td><strong>PSQI components</strong></td>
<td></td>
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<tr>
<td>1. Subjective sleep quality</td>
<td>1.57 (0.96)</td>
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<tr>
<td>2. Sleep latency</td>
<td>1.70 (1.08)</td>
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<td>3. Sleep duration</td>
<td>1.28 (1.14)</td>
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<td>4. Sleep efficiency</td>
<td>1.26 (1.23)</td>
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<tr>
<td>5. Sleep disturbance</td>
<td>1.96 (0.80)</td>
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<tr>
<td>6. Use of sleep medication</td>
<td>1.17 (1.34)</td>
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<tr>
<td>7. Daytime dysfunction</td>
<td>1.22 (0.84)</td>
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<tr>
<td><strong>PSQI items</strong></td>
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<tr>
<td>Average sleep duration (hr)</td>
<td>5.83 (1.77)</td>
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<tr>
<td>Habitual sleep efficiency (%)</td>
<td>75.00 (20.10)</td>
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<tr>
<td><strong>Depression</strong></td>
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<tr>
<td>CESD-R</td>
<td>17.19 (12.37)</td>
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<tr>
<td>CESD-R (minus sleep items)</td>
<td>14.10 (11.90)</td>
</tr>
<tr>
<td>Morbidity index</td>
<td>4.1 (3.1)</td>
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</table>

PSQI, Pittsburgh Sleep Quality Index; CESD-R, Center for Epidemiological Studies Depression Scale, Revised.
approached statistical significance (PSQI global sleep quality score >6 mean, 4.42; SD, 3.34; PSQI score ≤6 mean, 3.11; SD, 2.12; P = .06).

Sleep Disturbance and Race

Binary logistic regression models predicting sleep disturbance (PSQI global sleep quality score >6) are presented in Table 3. Black race was related to the presence of sleep disturbance (OR, 3.00; 95% CI, 1.17–7.69) in the unadjusted model (model 1). In subsequent models that adjusted for sex, employment status, income category (≥$20,000 or <$20,000 per year), education, CESD-R quartile, and chronic disease morbidity, race remained associated with global sleep quality (models 2 to 7). Only 2 predictors explained [my]10% of the association between race and sleep disturbance: income (attenuation in OR = [3.0 − 2.79] / [3.0 − 1] = 0.105) explained roughly 10.5% and education ([3.0 − 2.31] / [3.0 − 1] = 0.345) explained 34.5% of the risk associated with black race. Including all significant covariates in a model simultaneously (full model) yielded an association between race and sleep disturbance that was comparable in magnitude to the unadjusted association (OR, 2.44; 95% CI, 0.85–7.01). This same pattern of findings was observed in models using sleep duration (<6.5 hours per night) as the dependent variable, although race was not a significant predictor in models using sleep efficiency (<85%) as the dependent variable (data not shown).

Discussion

There was a high rate of sleep disturbance in this urban primary care sample (>70%). The mean global PSQI score in the sample was higher than that typically observed in samples of sleep apnea patients and was similar to the level of severity observed among patients suffering from primary insomnia, hemodialysis patients, and patients with lung cancer. Most notably, race was a significant independent correlate of sleep disturbance. This finding from a clinical sample is similar to that observed by Hall et al in a recent multisite community study of midlife women.

In the current study, being black was related to roughly 3 times the odds of having a sleep disturbance compared with being white/other. Race remained a significant correlate of sleep disturbance after controlling for income, employment, and a number of other factors, which explained approximately ≤10% of the association. Controlling for education, however, explained roughly 35% of the risk associated with race. Further longitudinal work might examine the biopsychosocial mechanisms causing these findings. The role of chronic stress and unmeasured lifestyle factors such as diet and exercise will be important to consider.

Limitations of this study include its cross-sectional nature, which prohibits any causal interpretations of these findings; the relatively small sample size; and the collection of data from only one primary care setting, which is not representative of the general population and may not represent other urban community samples. In addition, neither medication status nor body mass index were available in a manner that could be subjected to rigorous analysis, though both can contribute to sleep disturbance. Groups did not differ with respect to the PSQI component that measures sleep medications and several correlates of obesity (weight that was available from chart review, presence of diabetes taken from the chronic condition checklist, and presence of snoring derived from PSQI item 5e). It is also possible that the generally low educational level of this cohort may have introduced random error in responses on self-report instruments. This would make it less, rather than more, likely to detect associations. Finally, our overall sample was generally low on all socioeconomic indicators and contained mostly women, which limited the generalizability of the findings.

Notwithstanding these limitations, the strengths of the study include the use of a validated sleep instrument in an ethnically diverse sample to assess the contributions of race to sleep disturbance severity while adjusting for levels of depression and disease burden. The findings support the notion that a disparity exists for blacks with respect to sleep disturbances. Such results indicate a need for large-scale prevalence studies to determine whether this estimate is broadly generalizable. It will be important to test whether higher chronic stress is a key contributor to higher rates of poor sleep among blacks.

Further work is needed about whether midlife and older black patients in urban primary care settings may benefit from screening and/or increased monitoring for sleep problems.

Screening and monitoring is important for several reasons. First, sleep disturbance is becoming increasingly recognized as a risk factor for signifi-
Table 3. Percentages of Selected Characteristics by Sleep Disturbance* and Odds Ratios and Confidence Intervals from Hierarchical Logistic Regression Models Predicting Sleep Disturbance (n = 92)

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>PSQI score &gt;6 (%)</th>
<th>Model 1</th>
<th>Model 2</th>
<th>Model 3</th>
<th>Model 4</th>
<th>Model 5</th>
<th>Model 6</th>
<th>Model 7</th>
<th>Full Model</th>
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<td>Race/ethnicity</td>
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<tr>
<td>Black</td>
<td>81.3</td>
<td>3.00† (1.17–7.69)</td>
<td>2.97† (1.15–7.65)</td>
<td>2.92† (1.15–7.86)</td>
<td>2.71† (1.04–7.06)</td>
<td>2.39† (0.89–6.42)</td>
<td>2.91† (1.07–7.94)</td>
<td>3.15† (1.20–8.26)</td>
<td>2.44† (0.85–7.01)</td>
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<td>White/other</td>
<td>59.1</td>
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<tr>
<td>Male</td>
<td>66.7</td>
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<td>Female</td>
<td>71.8</td>
<td>1.13 (0.38–3.33)</td>
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<td>Employment status</td>
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<td>Employed (full or part time)</td>
<td>23.1</td>
<td>(referent)</td>
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<tr>
<td>Unemployed/retired/receiving disability</td>
<td>76.9</td>
<td>2.20 (0.82–5.92)</td>
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<td>Income per year</td>
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<td>&lt;$20,000</td>
<td>77.2</td>
<td>1.93 (0.75–4.97)</td>
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<td>≥$20,000</td>
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<td>No high school diploma or GED</td>
<td>83.3</td>
<td>3.51† (0.98–12.59)</td>
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<tr>
<td>High school diploma or GED and some college</td>
<td>82.8</td>
<td>3.63† (1.12–11.73)</td>
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<td>Associate’s degree or higher</td>
<td>53.8</td>
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<td>CESD-R§ quartiles</td>
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<td>Lowest quartile</td>
<td>45.0</td>
<td>(referent)†</td>
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<td>Second quartile</td>
<td>65.2</td>
<td>2.12 (0.60 to 7.56)</td>
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<td>Third quartile</td>
<td>80.8</td>
<td>4.28* (1.10 to 16.60)</td>
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<tr>
<td>Highest quartile</td>
<td>87.0</td>
<td>8.38* (1.78 to 39.60)</td>
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<td>Morbidity index**</td>
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<td>1.19* (0.99 to 1.42)</td>
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</table>

Model 1 provides the odds ratio (OR) (95% CI) for the unadjusted model of the association between race/ethnicity and sleep disturbance. Models 2 through 7 provide the OR (95% CI) for the adjusted models of the association between race/ethnicity and sleep disturbance (each model’s OR is adjusted for a single covariate). The Full Model provides the OR for the fully adjusted model of the association between race/ethnicity and sleep disturbance (after adjusting for all covariates).

*Pittsburgh Sleep Quality Index (PSQI) score >6.
†Wald statistic, \( P < .05 \).
‡Wald statistic, \( P < .1 \).
§Center for Epidemiological Studies Depression Scale, Revised (CESD-R) scores exclude sleep items.
¶The OR for education level represents the odds compared with the highest education level (having an Associate’s or higher degree).
‖The OR for each CESD-R quartile represents the odds compared to the lowest quartile.
**The morbidity index is a continuous variable.
cant morbidity and, to some extent, mortality. With respect to blacks in particular, further work is needed to test the extent to which sleep disturbances independently contribute to the established higher prevalence rates of hypertension and cardiovascular disease among these populations. Second, most common sleep disturbances can be effectively treated, but they remain vastly under-recognized and under-treated.13,48 Efficacious behavioral sleep medicine interventions, which could be integrated into primary care settings, exist for a variety of sleep disturbances.49 Finally, although the general population has poor access to such specialty services, it is likely that access is worse in traditionally underserved populations. Cost-effective means to deliver interventions for sleep disturbances in the primary care context merit consideration.

References

1. Geronimus AT, Hicken M, Keene D, Bound J. “Weathering” and age patterns of allostatic load scores among blacks and whites in the United States. Am J Public Health 2006;96:826–33.
27. Redline S, Tishler PV, Hans MG, Tosteson TD, Strohl K. Racial differences in sleep-disor-


