Religious Attendance: More Cost-Effective Than Lipitor?

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**Background:** A recent meta-analysis demonstrates a robust but small association between weekly religious attendance and longer life. However, the practical significance of this finding remains controversial.

**Methods:** Age-specific, actuarial death rates were modified according to published odds ratios to model the additional years of life attributable to: (1) weekly religious attendance; (2) regular physical exercise; and (3) statin-type lipid-lowering agents. Secondary analyses estimated the approximate cost for each additional year of life gained.

**Results:** Weekly attendance at religious services accounts for an additional 2 to 3 life-years compared with 3 to 5 life-years for physical exercise and 2.5 to 3.5 life-years for statin-type agents. The approximate cost per life-year gained was between $2,000 and $6,000 for regular exercise, $3,000 and $10,000 for regular religious attendance, and between $4,000 and $14,000 for statin-type agents.

**Conclusion:** The real-world, practical significance of regular religious attendance is comparable to commonly recommended therapies, and rough estimates even suggest that religious attendance may be more cost-effective than statins. Religious attendance is not a mode of medical therapy, but these findings warrant more and better quality research designed to examine the associations between religion and health, and the potential relevance such associations might have for medical practice. (J Am Board Fam Med 2006;19:103–9.)
all-cause, age-specific mortality rate according to the OR, it is possible to calculate and compare the life tables for the intervention and control arms of any given study. This article examines the practical importance of weekly religious attendance by comparing the improvement in life expectancy attributable to weekly religious attendance, regular physical exercise, and statin-type lipid-lowering agents.

This analysis is intended primarily as a thought experiment designed to explore the controversial association between religious attendance and longer life, and it is not intended for use in economic or clinical decision making. Although religious attendance appears to be a “health behavior” like physical exercise, it is not open to conventional modes of therapeutic manipulation for both ethical and scientific reasons. However, even if it is not therapeutically manipulable, there may be compelling reasons to study the association between religious attendance and mortality if the magnitude of that association is of sufficient practical importance.

**Methods**

Microsoft Excel (2003) was programmed to calculate life expectancy tables from age-specific mortality rates ($nM_x$) according to the method used by the National Center for Health Statistics. The probability of dying ($q_x$) between ages $x$ to $x + n$ was calculated such that $q_x = \frac{2 \cdot nM_x}{2 + nM_x}$ where $n$ is the number of years in the given age interval (in this case 5 years). The probability of death was constrained to 1.0 for ages over 85 years. The number of persons surviving to age $x$ ($l_x$) was calculated assuming a cohort of 100,000 live births such that $l_{x+n} = l_x - nd_x$ where $nd_x$ is the number of persons dying between ages $x$ to $x + n$ and is calculated by multiplying the number of persons alive at the beginning of the age interval ($l_x$) by the probability of dying within that age interval ($q_x$). The number of person-years lived between ages $x$ to $x + n$ ($L_x$) was calculated such that $L_x = (n/2)(l_x + l_{x+n})$. The total number of person-years lived above age $x$ ($T_x$) is the sum of the person-years lived to that age ($L_x$) and dividing the number of person-years lived above age $x$ ($T_x$) by the number of persons alive at age $x$ ($l_x$): $e_x = T_x / L_x$.

The spreadsheet programming was validated by entering the 2001 age-specific, all-cause mortality rates ($nM_x$) from survey data and comparing the calculated life expectancy with the published life tables from the same data. The spreadsheet accurately calculated life tables with values equivalent to the published tables.

The spreadsheet was then further programmed to calculate the improvement of life expectancy attributable to a given “modality” by adjusting the age-specific, all-cause mortality rate ($nM_x$) according to the published OR or relative risk (RR) for a given modality. Because ORs can be interpreted as relative risk estimates for studies when the incidence of the studied event is rare (<10%), and because the mortality rates for 2001 ranged from 0.02% to 7.12% the adjusted life tables were calculated using ORs as estimates of RR. Two life tables were calculated for each modality corresponding to the probabilities that constitute the numerator and denominator of relative risk. The first table described that portion of the population that accrued the benefit of the modality where $nM_x(adjusted) = nM_x(1 - (1 - RR)/(1 + RR))$. The second table described the life expectancy of that portion of the population that did not accrue the benefit of the modality where $nM_x(adjusted) = nM_x(1 + (1 - RR)/(1 + RR))$. The improvement of life expectancy ($e_x(0)$) attributable to a given “modality” was calculated for each age interval by subtracting the shorter life expectancy ($e_x(0)$) from the longer life expectancy ($e_x(l)$). Finally, the raw cost per life-year gained was calculated using an estimate of the yearly cost of the modality, multiplying the cost by the improvement of life expectancy ($e_x(l)$) and dividing it by the life years gained.

Using parameter estimates from published meta-analyses or reviews (see Table 1), comparative life tables were calculated for weekly religious attendance, regular physical exercise, and statin-type cholesterol-lowering drugs. Rough cost estimates for each modality focused exclusively on monetary costs without attempting to quantify other relevant costs such as time costs. Although individuals can exercise and attend religious services for free, this analysis assumes some modest financial commitment from those who exercise and attend religious services regularly. The average yearly cost for statin-type therapy ($836$) was derived from a published review of usage patterns of 6 statin brands in 1079 managed care plans.
cost of regular exercise was estimated as the annual cost of membership at the local gym ($500). The yearly cost of regular religious attendance was estimated by dividing the average yearly household contribution to religious institutions ($1336)\textsuperscript{14} by the average household size (2.59).\textsuperscript{15}

Sensitivity analyses were performed by calculating separate life tables for the upper and lower confidence intervals of the parameter estimates. An upper cost estimate of weekly religious attendance was also tested corresponding to the Old Testament “tithe” that required 10% of all income be offered to God. Published estimates of median household income ($42,000\textsuperscript{15}) and household size (2.59\textsuperscript{15}) yield a median, per capita tithe of $1,620.

Results

Table 2 summarizes the analysis of the 3 modalities. The first column records the standard life table for the 2001 all cause mortality data showing a 77.2 year life expectancy at birth. For each additional age range, the life expectancy is shown from the beginning of that age range. For example, according to the standard 2001 life table, a 40-year-old person is expected to live an additional 39.3 years.

Table 1. Published Estimates for Modeled Parameters

<table>
<thead>
<tr>
<th>Modality</th>
<th>OR/RR</th>
<th>Range*</th>
<th>Cost</th>
</tr>
</thead>
<tbody>
<tr>
<td>Regular physical exercise†</td>
<td>0.66\textsuperscript{11}</td>
<td>Not reported</td>
<td>$500</td>
</tr>
<tr>
<td>Statin-type medication</td>
<td>0.74\textsuperscript{12}</td>
<td>0.60 to 0.92</td>
<td>$836\textsuperscript{13}</td>
</tr>
<tr>
<td>Weekly religious attendance</td>
<td>0.775\textsuperscript{2}</td>
<td>0.719 to 0.833</td>
<td>$516\textsuperscript{14,15}</td>
</tr>
</tbody>
</table>

* Range corresponds to the 95% CI for each point estimate.
† Relative risk of regular physical exercise was calculated and reported separately by these authors for both women (0.66) and men (0.65). Given that these values are nearly identical, the more conservative estimate is used in this analysis.

Table 2. Comparative Life Tables

| Age Range | e\textsubscript{x(all cause)*} | e\textsubscript{a(l)} | e\textsubscript{a(s)} | e\textsubscript{a(g)} | $/e\textsubscript{a(g)} | e\textsubscript{a(l)} | e\textsubscript{a(s)} | e\textsubscript{a(g)} | $/e\textsubscript{a(g)} | e\textsubscript{a(l)} | e\textsubscript{a(s)} | e\textsubscript{a(g)} | $/e\textsubscript{a(g)} |
|-----------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|
| 0         | 77.2           | 80.1           | 75.0           | 5.2            | 7,772         | 79.3           | 75.5           | 3.7            | 17,816         | 78.9           | 75.8           | 3.1            | 12,950         |
| 1 to 4    | 76.7           | 79.7           | 74.5           | 5.2            | 7,675         | 78.8           | 75.1           | 3.7            | 17,592         | 78.5           | 75.3           | 3.2            | 12,787         |
| 5 to 9    | 72.8           | 75.8           | 70.6           | 5.2            | 7,318         | 74.9           | 71.2           | 3.7            | 16,766         | 74.5           | 71.4           | 3.2            | 12,184         |
| 10 to 14  | 67.9           | 70.8           | 65.6           | 5.2            | 6,850         | 69.9           | 66.2           | 3.7            | 15,679         | 69.6           | 66.4           | 3.2            | 11,392         |
| 15 to 19  | 62.9           | 65.9           | 60.7           | 5.2            | 6,382         | 65.0           | 61.3           | 3.7            | 14,593         | 64.6           | 61.5           | 3.1            | 10,600         |
| 20 to 24  | 58.1           | 61.0           | 55.9           | 5.1            | 5,944         | 60.2           | 56.5           | 3.7            | 13,581         | 59.8           | 56.7           | 3.1            | 9,860          |
| 25 to 29  | 53.4           | 56.3           | 51.2           | 5.0            | 5,571         | 55.4           | 51.8           | 3.6            | 12,718         | 55.1           | 52.0           | 3.1            | 9,231          |
| 30 to 34  | 48.7           | 51.5           | 46.5           | 5.0            | 5,175         | 50.6           | 47.0           | 3.6            | 11,801         | 50.3           | 47.3           | 3.0            | 8,563          |
| 35 to 39  | 43.9           | 46.7           | 41.8           | 4.9            | 4,771         | 45.8           | 42.3           | 3.5            | 10,868         | 45.5           | 42.5           | 3.0            | 7,881          |
| 40 to 44  | 39.3           | 42.0           | 37.2           | 4.8            | 4,382         | 41.2           | 37.7           | 3.5            | 9,965          | 40.8           | 37.9           | 2.9            | 7,223          |
| 45 to 49  | 34.7           | 37.4           | 32.7           | 4.7            | 4,006         | 36.6           | 33.2           | 3.4            | 8,904          | 36.2           | 33.4           | 2.8            | 6,587          |
| 50 to 54  | 30.3           | 32.9           | 28.3           | 4.5            | 3,645         | 32.1           | 28.8           | 3.2            | 7,716          | 31.8           | 29.0           | 2.7            | 5,978          |
| 55 to 59  | 26.0           | 28.5           | 24.2           | 4.3            | 3,295         | 27.7           | 24.6           | 3.1            | 7,452          | 27.4           | 24.8           | 2.6            | 5,388          |
| 60 to 64  | 21.9           | 24.3           | 20.2           | 4.1            | 2,962         | 23.6           | 20.6           | 2.9            | 6,680          | 23.3           | 20.8           | 2.5            | 4,825          |
| 65 to 69  | 18.1           | 20.4           | 16.5           | 3.8            | 2,655         | 19.7           | 16.9           | 2.8            | 5,968          | 19.4           | 17.1           | 2.3            | 4,306          |
| 70 to 74  | 14.7           | 16.7           | 13.2           | 3.5            | 2,365         | 16.1           | 13.6           | 2.5            | 5,297          | 15.9           | 13.7           | 2.1            | 3,815          |
| 75 to 79  | 11.6           | 13.5           | 10.2           | 3.2            | 2,081         | 12.9           | 10.6           | 2.3            | 4,639          | 12.6           | 10.7           | 2.0            | 3,334          |
| 80 to 84  | 8.9            | 10.6           | 7.7            | 3.0            | 1,790         | 10.1           | 7.9            | 2.1            | 3,962          | 9.9            | 8.1            | 1.8            | 2,840          |
| 85+       | 6.6            | 8.3            | 5.5            | 2.8            | 1,471         | 7.8            | 5.8            | 2.0            | 3,215          | 7.6            | 5.9            | 1.7            | 2,293          |

* e\textsubscript{x(all cause)} is the standard all cause life expectancy as published in the census data.\textsuperscript{9} e\textsubscript{a(l)} is the longer life expectancy attributable to the modality’s health benefit. e\textsubscript{a(s)} is the shorter life expectancy predicted for those not exposed to the modality’s health benefit. e\textsubscript{a(g)} is the years of life gained attributable to the specific modality. $/e\textsubscript{a(g)} is the raw cost per life year gained through exposure to the specific modality. Rows for ages 20, 40, and 70 are emphasized for ease of comparison.
for a total life span of 79.3 years. Alternatively, an
85-year-old is expected to live only 6.6 more years.

The remaining columns show the life tables ad-
justed to correspond with the 3 modeled modalities
of statin-type therapy for high-risk coronary artery
disease (CAD), regular physical exercise and weekly
religious attendance. Consider the first columns
corresponding to regular physical exercise. A 40-
year-old person who exercises regularly is expected
to live 42.0 years for a total life span of 82.0 years
whereas a similar 40-year-old who does not exer-
rise is expected to live only 37.2 years. For such a
40-year-old person, regular exercise accounts for
an extra 4.8 years of life. If the cost of such regular
exercise were equivalent to a membership in a local
gym, the raw cost per year of life gained through
regular exercise would be $4,382. The amount of
life gained and the cost per year of life gained vary
throughout the life table depending on the age at
which an individual starts exercising regularly
(thereby accruing the health benefit). For example,
regular exercise from the age of 20 accrues 5.1
additional years of life at a cost of $5,944 per year
of life gained. Alternatively, regular exercise from
the age of 70 accrues only 3.5 years of additional
life at a cost of $2,465 per year of life gained.

In similar fashion, a 40-year-old person with
CAD taking statins is likely to live 3.5 years longer
than a similar person not taking statins, and the raw
cost per year of life gained for initiating statin
therapy at age 40 is $9,965. Likewise, a 40-year-old
person who attends religious services regularly is
predicted to live 2.9 years longer than a similar
person who does not attend religious services, and
the raw cost per year of life gained for attending
religious services from the age of 40 is $7,223.

Life expectancies and costs are reported for all
age ranges and can be interpreted from Table 2 in
a similar fashion. However, the adjusted life expect-
cancies for each modality from birth to age 20 are	abulated only for the sake of completeness. It
makes little sense to suppose that an infant or
young adult would have CAD requiring statin ther-
apy. Likewise, none of the data examining religious
attendance or regular exercise included subjects
younger than 20 years of age, and it is therefore not
possible to draw any justifiable conclusions from
the first 5 rows of the table.

Full life tables for the sensitivity analyses are not
shown, but estimates from the 60-year-old age
range of the relevant life table are reported below.

Age 60 was chosen because the average age of
subjects in the meta-analysis of religious attendance
was approximately 60. At the upper confidence
interval, statins and religious attendance accounted
for 5.1 ($4,104) and 3.2 ($3,785) additional years of
life (and cost per year of life gained), respectively.

At the lower confidence interval, statins and reli-
gious attendance accounted for 0.8 ($23,042) and
1.8 ($6,631) additional years of life, respectively.

No sensitivity analysis for physical exercise was
possible because no confidence intervals were re-
ported. Finally, assuming a complete tithe of
household income, the upper limit for the cost of
religious attendance per year of life gained ranged
from $15,000 if the tithe began at age 60 to $31,000
if the tithe began at age 20.

Discussion

Regular physical exercise, statin-type therapy and
weekly religious attendance account for an addi-
tional 2 to 5 years of life, suggesting that the real
world, practical significance of weekly religious at-
tendance is of similar magnitude as these other
widely recommended therapies or health behaviors.
Regular physical exercise was both the most effec-
tive (3 to 5 additional years of life) and least expen-
sive ($2,000 to $6,000 per year of life gained).
Statins were more effective than regular religious
attendance (2.1 to 3.7 vs 1.8 to 3.1 additional years
of life), but they were also more expensive ($4,000
to $14,000 vs. $3,000 to $10,000 per year of life
gained).

Sensitivity analyses suggest that the lowest esti-
mate for the years of life gained attributable to
regular religious attendance is 1.8. The practical
significance of nearly 2 additional years of life is
intuitively obvious, but it is also nearly twice as
large as the estimated health liability attributable to
obesity that led Olshansky et al16 to conclude that
obesity may lead to an overall decline in life expect-
cy in the United States in the 21st century.

Although the secondary analysis of cost makes
no claims to be a formal cost-benefit analysis, the
estimate calculated here ($4,000 to $23,000 per
life-year gained) corresponds to previously pub-
lished cost-effectiveness ratios (CERs) for statin-
based therapy that range from $5,000 to $50,000
per year of life gained. The cost per life-year
 gained attributable to religious attendance is as
high as $31,000, but even this estimate falls within
the range of CERs that previously published literature judges to be societally acceptable. Thresholds have been proposed between $20,000 and $100,000 per life-year gained, with most accepted modalities falling below $40,000.\textsuperscript{20} However, many commonly accepted medical interventions have significantly higher CERs such as hemodialysis\textsuperscript{21} ($274,000) or cardiopulmonary resuscitation\textsuperscript{22} ($226,000).

Within the admittedly limited assumptions of this thought experiment, religious attendance appears to be more cost-effective than statin-type lipid-lowering agents. However, this analysis should not be interpreted to mean that health care payers should start covering the annual tithe of religious patients. Religious attendance is not a therapy, and although the association between attendance and health is well established, there is no evidence that changing attendance causes a change in health outcomes. (The data supporting the association are derived from prospective, observation epidemiologic studies.) Furthermore, even if a patient did start attending church to get well, there are ethical, theological, and methodological problems with this instrumental approach to the associations between faith and health.\textsuperscript{6,7} For example, it is not clear that the observed reduction in morality would accrue due to religious attendance. From a theological perspective such instrumental use of religion is idolatrous. From a methodological perspective, it is not at all clear that “instrumental faith” is sufficiently genuine to accrue the observed reduction in mortality. And finally, neither of these qualifications address the ethical quagmire engendered by any medical recommendation to attend religious services.

However, given the fact that religious attendance might account for up to 3 additional years of life, the associations between religiousness and health may remain relevant for medical practice even if such religiousness is not therapeutically manipulable. Rather than approaching religious belief and practice as a health behavior, it may be more appropriate to approach religiousness as a demographic factor.\textsuperscript{6} For example, the incidence of gastric cancer is higher among Japanese men, and knowledge of this fact might guide a physician to initiate early and frequent screening for gastric cancer among male Japanese patients. It is not possible to manipulate or therapeutically change a patient’s ethnicity, but knowledge of the association between ethnicity and gastric cancer remains relevant for appropriate medical care. Likewise, even though it is not therapeutically manipulable, specific religious beliefs and practices may be associated with both positive and negative health outcomes that might appropriately influence medical care and decision-making. The details of the associations between religiousness and health are far from clear, but the analysis described here suggests that the practical significance of the association between mortality and religious attendance is sufficiently large to warrant more and better quality research designed to examine the associations between religion and health, and the potential relevance such associations might have for medical practice.

As with any analysis of data, this study has several limitations. First, the data regarding statin-type lipid-lowering agents are derived from randomized controlled trials (RCTs), and are therefore a stronger form of evidence than the epidemiologic cohort studies that describe the association between religious attendance and mortality. It is also not entirely appropriate to compare a secondary prevention therapy like statins to religious attendance that, if anything, is most analogous to primary prevention. Furthermore, the point estimate used for the effectiveness of statin-type therapy is applicable only to populations with high-risk CAD, which corresponds to only 0.7% of 50- to 59-year-old men and as few as 0.1% of 40- to 49-year-old men.\textsuperscript{12} While acknowledging this asymmetry, this study included statin-type therapy as a comparator for 2 reasons. First, statin-type therapy is frequently used as a quasi-primary prevention of CAD,\textsuperscript{19} and second, it provides continuity with the previously published debate regarding the relevance of weekly religious attendance.\textsuperscript{2,3}

To make a more direct comparison, the study also analyzed the mortality data regarding regular physical exercise. Despite the widely accepted practice of encouraging regular physical exercise, there are no RCTs investigating the impact of regular exercise on mortality in healthy populations. Rather, the data supporting the benefits of regular exercise derive from large epidemiologic studies that are not categorically different in design from the studies examining religious attendance. In fact, the epidemiologic findings regarding regular exercise are so widely accepted that nobody has actually invested the effort to perform a rigorous meta-
analysis of the health benefits of regular exercise. After a careful search of the literature and consultation with experts in the field, this analysis used the largest and most complete quantitative review of existing data regarding regular exercise.11

Although the meta-analysis of religious attendance is actually a stronger form of evidence than the quantitative review of regular exercise, it is not without its own limitations. McCullough et al11 note that their analysis did not control for all relevant confounding variables. However, the 2 best controlled, single-population studies demonstrate similar or more favorable hazard ratios than those modeled here (0.7723 and 0.6724). In fact, critics of McCullough’s analysis concede that it is methodologically rigorous, comparable in quality to other meta-analyses, and definitive; but their critique challenges the meaning of the finding rather than impeaching the finding itself. Indeed, the same critics have recently published their own well-controlled analysis of the EPESE cohorts that found a relative risk of religious attendance (RR = 0.78, 95% CI, 0.70, 0.88) nearly identical to the risk modeled here.25

Finally, one potential source of bias in any study is the investigator’s world view (or paradigm) that constrains both the types of questions that can be asked and the available answers. Such paradigm bias may be especially relevant in the study of religious belief and practice. Although the author of this study is a practicing Christian, he has worked to ensure that the design and interpretation of this study can stand on purely scientific grounds.

Despite the limitations of the available data, this analysis suggests that the practical, real world significance of regular religious attendance (2 to 3 additional years of life) is similar to widely recommended health practices like regular physical exercise and statin-based therapy of CAD. Rather than dismissing this finding as weak or nonexistent,3 it may be more fruitful to invest the necessary resources to better understand the nature and relevance of the associations between religious attendance and health.

References


